

Linear Regression ¹

Shaobo Li

University of Kansas

¹Partially based on Hastie, et al. (2009) ESL, and James, et al. (2013) ISLR

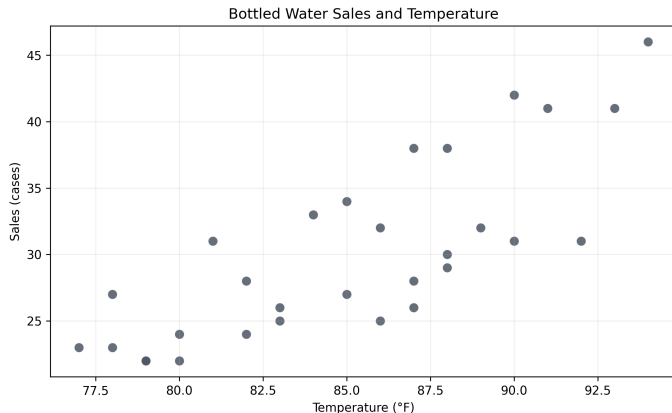
Linear regression – a fundamental learning algorithm

- Supervised learning method
- It assumes the dependence of Y on X is linear
- Largely used in many disciplines
- Simple and interpretable
- Fundamental in data science

What can linear regression do?

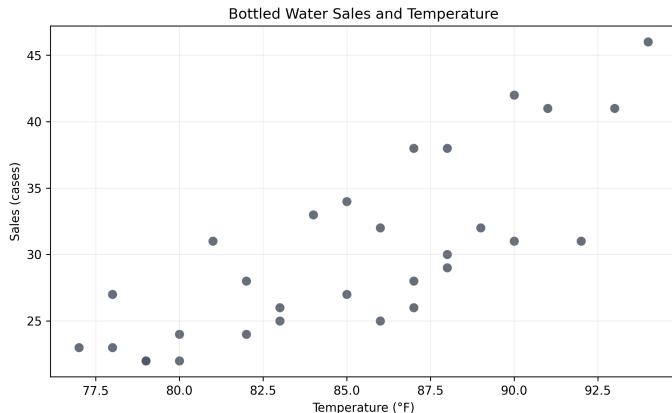
- Is there an association between X and Y ?
- If yes, how strong is this association?
- Is this association linear?
- If there are multiple X 's (X_1 , X_2 and X_3), which of them are related to Y and which are not?
- Can we predict the value of Y for any given X ?
- How accurate is such prediction?

An example – bottled water sales



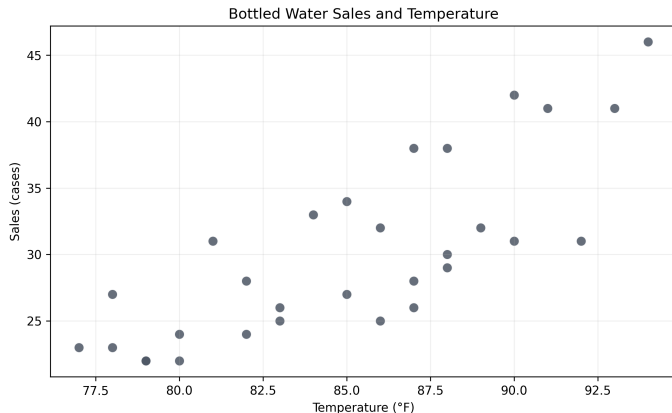
- Is there an association between *Temperature* and *Sales*?

An example – bottled water sales



- Is there an association between *Temperature* and *Sales*?
- If yes, how strong is this association?

An example – bottled water sales



- Is there an association between *Temperature* and *Sales*?
- If yes, how strong is this association?
- Is this association linear?

An example – bottled water sales

- We can write this relationship as

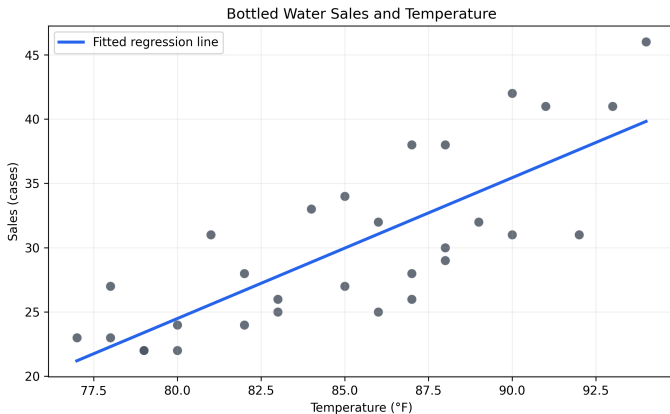
$$Sales \approx \beta_0 + \beta_1 \times Temperature$$

- We use “ $\approx$ ” because the model always approximates the “truth”.
- This is a *simple linear regression*.
- β_0 is called intercept and β_1 is slope.
- Given the data points we observed, the model is estimated to be

$$Sales \approx -63.11 + 1.09 \times Temperature$$

- How do we interpret this model?

Fitted straight line – a simple linear regression



Linear regression models

- More generally, a simple linear regression model is

$$Y = \beta_0 + \beta_1 X + \epsilon,$$

and a multiple linear regression model is

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \epsilon,$$

where ϵ catches the error of the “model” from the “truth”.

- Y is called dependent variable (or response, outcome).
- X is called independent variable (or covariates, explanatory variable).

Linear regression model in matrix form

$$\begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & X_{11} & X_{12} & \dots & X_{1p} \\ 1 & X_{21} & X_{22} & \dots & X_{2p} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & X_{n1} & X_{n2} & \dots & X_{np} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

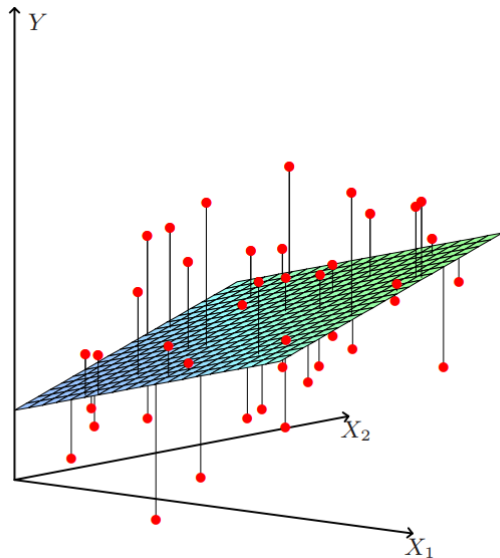
- $\mathbf{X}$ is called design matrix

- The *estimated linear regression model* is

$$\hat{y} = \mathbb{E}(y|\mathbf{X}) = \mathbf{X}\hat{\beta}$$

- We need to figure out $\hat{\beta}$, the estimates of β
- Method to use: *ordinary least square (OLS)*

Least square solution



Least square solution

- We want to minimize residual sum squares (RSS)

$$\begin{aligned}RSS(\beta) &= \sum_{i=1}^n (y_i - \mathbf{x}_i^T \beta)^2 \\ &= (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta)\end{aligned}$$

- Take first-order derivative with respect to β and set to 0

$$\begin{aligned}0 &= \frac{\partial RSS(\beta)}{\partial \beta} = -2\mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta) \\ \mathbf{X}^T \mathbf{y} &= \mathbf{X}^T \mathbf{X} \beta\end{aligned}$$

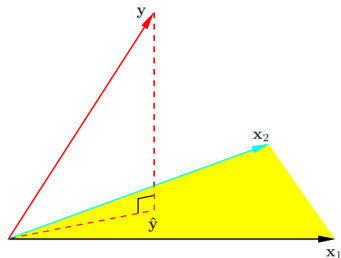
- This is called *normal equation*.

Least square solution

- By assuming $p < n$, the OLS solution is

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- The predicted value is $\hat{\mathbf{y}} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$
- $\mathbf{H} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ is called hat matrix or projection matrix
- That is, $\hat{\mathbf{y}} = \mathbf{H}\mathbf{y}$. In other words, $\hat{\mathbf{y}}$ is a linear projection of $\mathbf{y}$



Some important questions after fitting the model

- How to interpret the model?
- Is at least one of the predictors useful to predict and explain the response?
- Do all predictors help to explain response, or just a subset?
- How well does the model fit the data?
- Given a set of predictor values, what response value does the model predict? How accurate is the prediction?

Hypothesis testing – test multiple coefficients

Is at least one X useful?

- F -test for overall significance

- $H_0: \beta_1 = \dots = \beta_p = 0$; H_1 : at least one $\beta \neq 0$
- F statistics

$$F^* = \frac{(TSS - RSS)/p}{RSS/(n - p - 1)} \sim F_{p, n-p-1}$$

where $TSS = \sum_{i=1}^n (y_i - \bar{y})^2$, is total sum squares

Hypothesis testing – test individual coefficients

Is a specific X relevant?

- Testing for individual β
 - $H_0: \beta_j = 0$; $H_1: \beta_j \neq 0$
 - Using T-test since the true variance is unknown

$$T = \frac{\hat{\beta}_j}{se(\hat{\beta}_j)} = \frac{\hat{\beta}_j}{\hat{\sigma} \sqrt{v_j}} \sim t_{n-p-1}$$

where v_j is the j th diagonal element of $(\mathbf{X}^T \mathbf{X})^{-1}$

- Reject H_0 if p-value $< \alpha$ or $|T| > T_{1-\alpha}^{(n-p-1)}$
- Confidence interval: $\hat{\beta} \pm se(\hat{\beta}) \times T_{1-\alpha}^{(n-p-1)}$

R output for bottle water example

```
> model1 <- lm(Sales~Temperature, data = water)
> summary(model1)
```

Call:

```
lm(formula = Sales ~ Temperature, data = water)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-6.6253	-2.9128	-0.9378	4.1109	6.5647

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-63.1146	13.7045	-4.605	8.15e-05	***
Temperature	1.0950	0.1609	6.807	2.15e-07	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 4.145 on 28 degrees of freedom

Multiple R-squared: 0.6233, Adjusted R-squared: 0.6099

F-statistic: 46.34 on 1 and 28 DF, p-value: 2.148e-07

Python output for bottle water example

```
import statsmodels.formula.api as smf
temperature_model = smf.ols("Sales ~ Temperature", data=water).fit()
print(temperature_model.summary())
```

✓ 57ms 9 mins ago

OLS Regression Results

```
=====
```

Dep. Variable:	Sales	R-squared:	0.623
Model:	OLS	Adj. R-squared:	0.610
Method:	Least Squares	F-statistic:	46.34
Date:	Wed, 09 Sep 2026	Prob (F-statistic):	2.15e-07
Time:	13:34:29	Log-Likelihood:	-84.187
No. Observations:	30	AIC:	172.4
Df Residuals:	28	BIC:	175.2
Df Model:	1		
Covariance Type:	nonrobust		

```
=====
```

	coef	std err	t	P> t	[0.025	0.975]
Intercept	-63.1146	13.705	-4.605	0.000	-91.187	-35.042
Temperature	1.0950	0.161	6.807	0.000	0.765	1.425

```
=====
```

Omnibus:	6.685	Durbin-Watson:	0.983
Prob(Omnibus):	0.035	Jarque-Bera (JB):	2.036
Skew:	0.113	Prob(JB):	0.361
Kurtosis:	1.744	Cond. No.	1.54e+03

```
=====
```

Model Assessment – R Square and MSE

- It is proportion of variation in Y explained by the model

$$R^2 = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS}$$

R^2 increases monotonically as number of X 's increasing.

- Adjusted R^2

$$R_{adj}^2 = 1 - \frac{n-1}{n-p-1} \frac{RSS}{TSS}$$

- Mean Squared Error (MSE)

$$MSE = \frac{1}{n-p-1} \times RSS$$

It is an unbiased estimate of σ^2 , variance of ϵ . (Can you show this?)

Categorical covariate

Now consider another variable "Game Day".

$$\text{Game Day} = \begin{cases} 1 & \text{game day} \\ 0 & \text{non-game day} \end{cases}$$

- Sales (Y); Temperature (X_1); Game Day (X_2)
- How does the linear model look like?

Categorical covariate

Now consider another variable "Game Day".

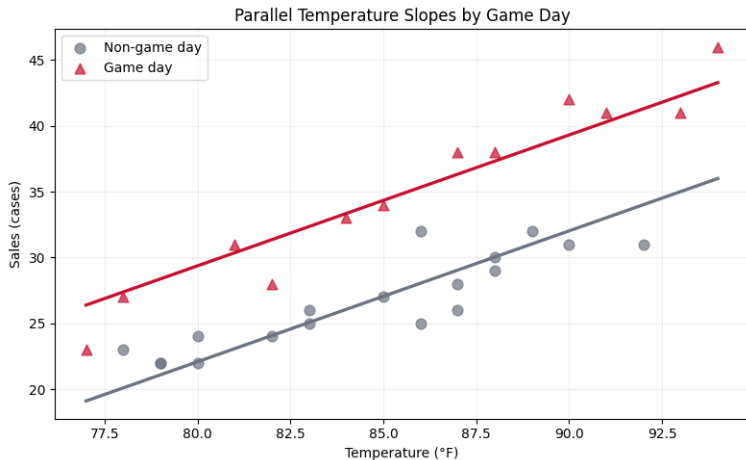
$$\text{Game Day} = \begin{cases} 1 & \text{game day} \\ 0 & \text{non-game day} \end{cases}$$

- Sales (Y); Temperature (X_1); Game Day (X_2)
- How does the linear model look like?

$$\begin{aligned} E(Y|X_1, X_2) &= \beta_0 + \beta_1 X_1 + \beta_2 X_2 \\ &= \begin{cases} \beta_0 + \beta_1 X_1 + \beta_2 & \text{for game day} \\ \beta_0 + \beta_1 X_1 & \text{for non-game day} \end{cases} \end{aligned}$$

- How do we interpret this model?
- Can you draw a graph to illustrate this model?

Parallel Fitted Lines



More than two categories

- Consider another covariate – *Location* (X_3), which has 3 categories: Campus, Downtown, and Stadium.
- How can we write the model?

More than two categories

- Consider another covariate – *Location* (X_3), which has 3 categories: Campus, Downtown, and Stadium.

- How can we write the model?

Following the same logic, we can write:

$$E(Y|X_1, X_2, X_3) = \begin{cases} \beta_0 + \beta_1 X_1 + \beta_2 X_2 & \text{for Campus} \\ \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_{3,1} & \text{for Downtown} \\ \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_{3,2} & \text{for Stadium} \end{cases}$$

- Can you see how do we handle the categorical variable *Location* (X_3)?
- How do we draw conclusions in testing the significant of X_3 ?

Python output

```

=====
                        OLS Regression Results
=====
Dep. Variable:          Sales    R-squared:                0.689
Model:                  OLS      Adj. R-squared:           0.653
Method:                 Least Squares    F-statistic:             19.22
Date:                   Wed, 09 Sep 2026    Prob (F-statistic):      8.92e-07
Time:                   17:11:06    Log-Likelihood:         -81.306
No. Observations:      30    AIC:                     170.6
Df Residuals:          26    BIC:                     176.2
Df Model:               3
Covariance Type:       nonrobust
=====
                        coef    std err          t      P>|t|      [0.025    0.975]
-----
Intercept             -60.0970     13.645     -4.404     0.000    -88.144    -32.050
C(Location)[T.Downtown]  1.5354     1.891     0.812     0.424     -2.351     5.422
C(Location)[T.Stadium]   4.1102     1.791     2.295     0.030     0.428     7.792
Temperature             1.0374     0.164     6.321     0.000     0.700     1.375
=====
Omnibus:               8.510    Durbin-Watson:           0.801
Prob(Omnibus):         0.014    Jarque-Bera (JB):        2.199
Skew:                  -0.023    Prob(JB):                0.333
Kurtosis:              1.674    Cond. No.                 1.63e+03
=====
```

Can you write the estimated model and interpret the coefficients?

- Consider the same example without Location.

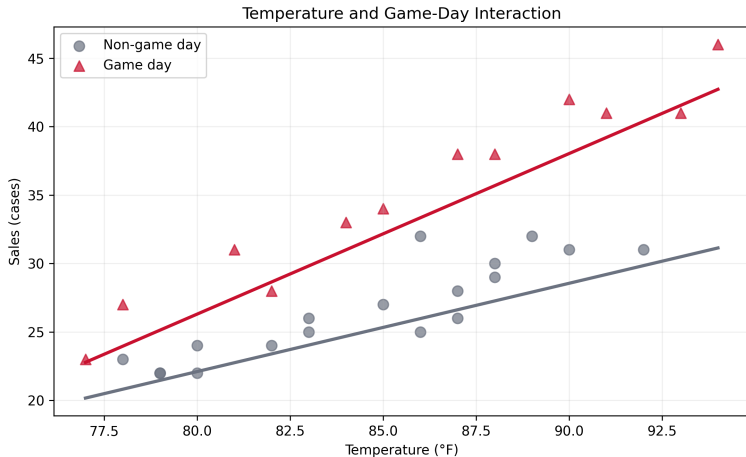
$$\begin{aligned} E(Y|X_1, X_2) &= \beta_0 + \beta_1 X_1 + \beta_2 X_2 \\ &= \begin{cases} \beta_0 + \beta_1 X_1 + \beta_2 & \text{for game day} \\ \beta_0 + \beta_1 X_1 & \text{for non-game day} \end{cases} \end{aligned}$$

- This model says Temperature (X_1) has the same effect regardless it is a game day or not. But what if the effects are different?
- To answer this question, we need to examine the **interaction effect**

$$E(Y|X_1, X_2) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 (X_1 * X_2)$$

- Can you break down this model with respect to GameDay?
- Can you draw a graph to illustrate this model?

Parallel Fitted Lines

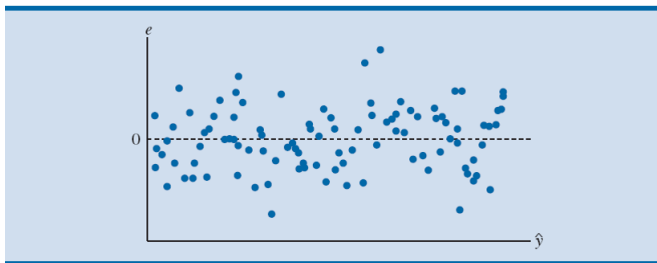


Model Diagnostics

■ Model assumptions:

- Linear relationships between Y and X 's
- The error term $\{\epsilon_1, \dots, \epsilon_n\} \stackrel{i.i.d.}{\sim} N(0, \sigma^2)$
Independent normal distribution; $\mathbb{E}(\epsilon_i) = 0$; $\text{Var}(\epsilon_i) = \text{constant}$.

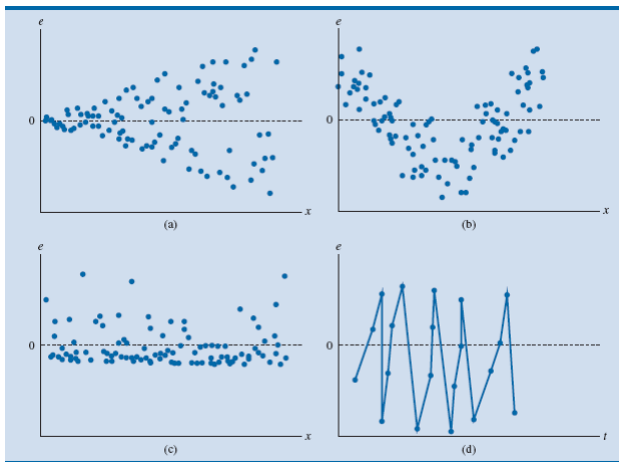
■ Residual plot (an ideal residual plot looks like this)²



²source: Camm, et al., *Essentials of Business Analytics*

Residual plot

Which type of assumption is violated?



3

³source: Camm, et al., *Essentials of Business Analytics*

Other Diagnostic Plots

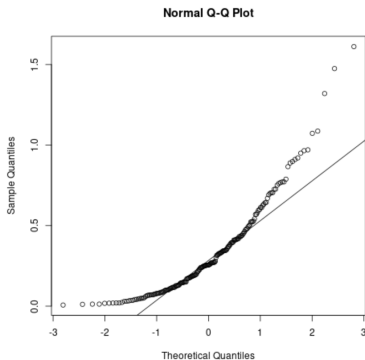
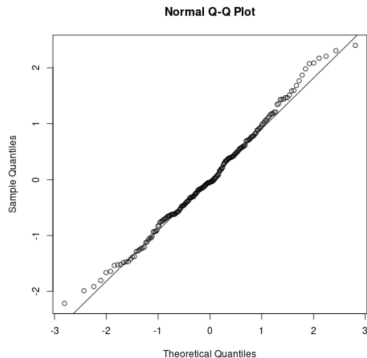
■ Normal Quantile-Quantile Plot

- It plots the standardized residual vs. theoretical quantiles
- An easy way to visually test the normality assumption
- If residual follows normal distribution, you should expect all dots lie on the diagonal straight line.

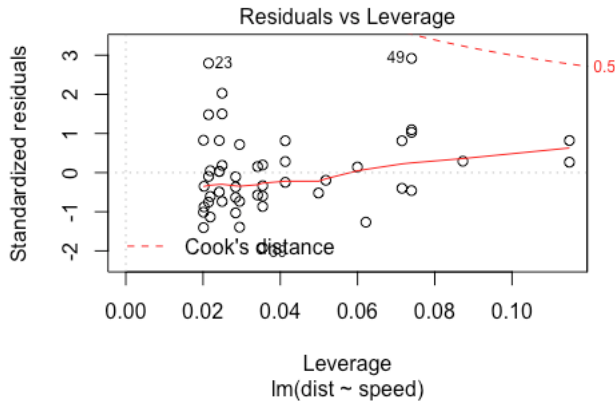
■ Residual-Leverage Plot

- This plot checks if there are any influential points, which could alter your analysis by excluding them
- The points that lie outside the dashed line, Cook's distance, are considered as influential points

Normal Q-Q plot



Residual leverage plot



Model Building Process

