Linear Regression ¹

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¹Partially based on Hastie, et al. (2009) ESL, and James, et al. (2013) ISLR

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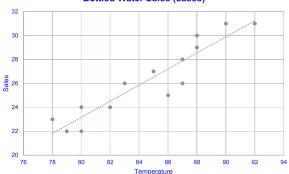
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- Supervised learning method
- It assumes the dependence of Y on X is linear
- Largely used in many disciplines
- Simple and interpretable
- Fundamental in data science

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- Is there an association between X and Y?
- If yes, how strong is this association?
- Is this association linear?
- If there are multiple X's (X₁, X₂ and X₃), which of them are related to Y and which are not?
- Can we predict the value of Y for any given X?
- How accurate is such prediction?

An example – bottled water sales

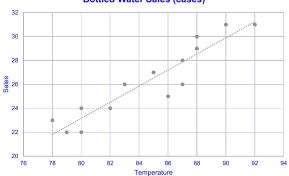


Bottled Water Sales (cases)

■ Is there an association between *Temperature* and *Sales*?

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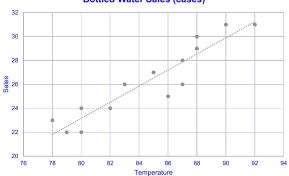


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Bottled Water Sales (cases)

- Is there an association between *Temperature* and *Sales*?
- If yes, how strong is this association?
- Is this association linear?

MS in Business Analytics

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We can write this relationship as

Sales $\approx \beta_0 + \beta_1 \times Temperature$

• We use " \approx " because model alway approximates the "truth".

- This is a *simple linear regression*.
- β_0 is called intercept and β_1 is slope.
- Given the data points we observed, the model is estimated to be

Sales $\approx -30.70 + 0.67 \times Temperature$

This is the straight line we saw before.

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More generally, a simple linear regression model is

$$Y = \beta_0 + \beta_1 X + \epsilon,$$

and a multiple linear regression model is

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \epsilon,$$

where ϵ catches the error between the "model" and the "truth".

- *Y* is called dependent variable (or response, outcome).
- X is called independent variable (or covariates, explanatory variable).

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Linear regression model in matrix form

$$\begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & X_{11} & X_{12} & \dots & X_{1p} \\ 1 & X_{21} & X_{22} & \dots & X_{2p} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & X_{n1} & X_{n2} & \dots & X_{np} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

X is called design matrix

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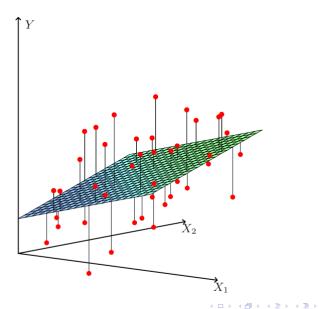
The estimated linear regression model is

$$\hat{oldsymbol{y}} = \mathbb{E}(oldsymbol{y}|oldsymbol{\mathsf{X}}) = oldsymbol{\mathsf{X}}\hat{oldsymbol{eta}}$$

- ${\scriptstyle \blacksquare}$ We need to figure out $\hat{\beta},$ the estimates of β
- Method to use: ordinary least square (OLS)

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Least square solution



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Least square solution

We want to minimize residual sum squares (RSS)

$$RSS(\beta) = \sum_{i=1}^{n} (y_i - \mathbf{x}_i^T \beta)^2$$
$$= (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta)$$

Take first-order derivative with respect to β and set to 0

$$0 = \frac{\partial RSS(\beta)}{\partial \beta} = -2\mathbf{X}^{T}(\mathbf{y} - \mathbf{X}\beta)$$
$$\mathbf{X}^{T}\mathbf{y} = \mathbf{X}^{T}\mathbf{X}\beta$$

This is called *normal equation*.

Exercise: derive the closed form of the solution

Least square solution

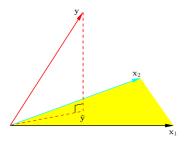
• By assuming p < n, the OLS solution is

 $\hat{oldsymbol{eta}} = (\mathbf{X}^{ op} \mathbf{X})^{-1} \mathbf{X}^{ op} \mathbf{y}$

• The predicted value is $\hat{\mathbf{y}} = \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$

• $\mathbf{H} = \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ is called hat matrix or projection matrix

That is, $\hat{\mathbf{y}} = \mathbf{H}\mathbf{y}$. In other words, $\hat{\mathbf{y}}$ is a linear projection of \mathbf{y}



- Is at least one of the predictors useful to predict and explain the response?
- Do all predictors help to explain response, or just a subset?
- How well does the model fit the data?
- Given a set of predictor values, what response value does the model predict? How accurate is the prediction?

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Is at least one X useful?

- F-test for overall significance
 - H_0 : $\beta_1 = \ldots = \beta_p = 0$; H_1 : at least one $\beta \neq 0$
 - F statistics

$$F^* = rac{(TSS - RSS)/p}{RSS/(n-p-1)} \sim F_{p,n-p-1}$$

where $TSS = \sum_{i=1}^{n} (y_i - \bar{y})^2$, is total sum squares

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Is a specific X relevant?

- \blacksquare Testing for individual β
 - $H_0: \ \beta_j = 0; \ H_1: \ \beta_j \neq 0$
 - Using T-test since the true variance is unknown

$$T = \frac{\hat{\beta}_j}{se(\hat{\beta}_j)} = \frac{\hat{\beta}_j}{\hat{\sigma}\sqrt{v_j}} \sim t_{n-p-1}$$

where v_j is the *j*th diagonal element of $(\mathbf{X}^T \mathbf{X})^{-1}$ - Reject H_0 if p-value $< \alpha$ or $|T| > T_{1-\alpha}^{(n-p-1)}$

• Confidence interval: $\hat{\beta} \pm se(\hat{\beta}) \times T_{1-\alpha}^{(n-p-1)}$

```
> model1<- lm(Sales~Temperature. data = sales)</pre>
> summarv(model1)
Call:
lm(formula = Sales ~ Temperature, data = sales)
Residuals:
   Min
            10 Median 30
                                  Max
-2.1994 -0.5016 0.2908 0.8350 1.4542
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -30.69720 6.38033 -4.811 0.000425 ***
Temperature 0.67322 0.07529 8.942 1.18e-06 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 1.184 on 12 degrees of freedom
Multiple R-squared: 0.8695. Adjusted R-squared: 0.8586
F-statistic: 79.96 on 1 and 12 DF, p-value: 1.182e-06
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- Resampling method
- A powerful tool to quantify uncertainty
 - standard error
 - confidence interval
- Random sampling with replacement
- The general procedure:
 - fit a model ${\boldsymbol{B}}$ times based on ${\boldsymbol{B}}$ bootstrap samples
 - store all the parameter estimates
 - calculate standard error and confidence interval

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Model Assessment – R Square and MSE

It is proportion of variation in Y explained by the model

$$R^2 = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS}$$

R² increases monotonically as number of X's increasing.Adjusted R²

$$R_{adj}^2 = 1 - \frac{n-1}{n-p-1} \frac{RSS}{TSS}$$

Mean Squared Error (MSE)

$$MSE = \frac{1}{n-p-1} \times RSS$$

It is an unbiased estimate of σ^2 , variance of ϵ . (Can you show this?)

Akaike information criterion (AIC), the smaller the better

$$AIC = -2\log(\hat{L}) + 2p$$

Bayesian information criteria (BIC), the smaller the better

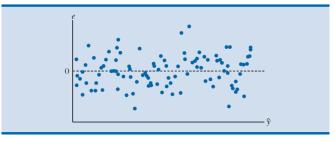
$$BIC = -2\log(\hat{L}) + \log(n)p$$

where \hat{L} is estimated likelihood function

- Mellow's C_p is the same as AIC for linear regression
- Cross-validation error (to assess prediction accuracy)
- These metrics are important for model selection

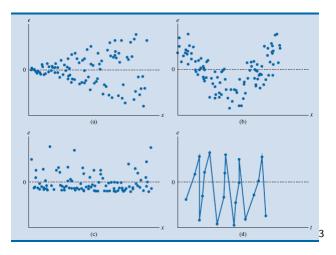
- Model assumptions:
 - Linear relationships between Y and X's
 - The error term $\{\epsilon_i, \ldots, \epsilon_n\} \stackrel{i.i.d.}{\sim} N(0, \sigma^2)$ Independent normal distribution; $\mathbb{E}(\epsilon_i) = 0$; $Var(\epsilon_i) = \text{constant}$.

Residual plot (an ideal residual plot looks like this)²



Residual plot

Which type of assumption is violated?



Normal Quantile-Quantile Plot

- It plots the standardized residual vs. theoretical quantiles
- An easy way to visually test the normality assumption
- If residual follows normal distribution, you should expect all dots lie on the diagonal straight line.

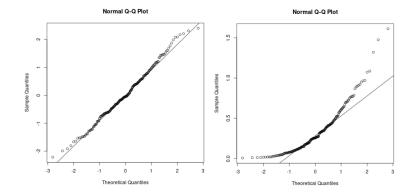
Residual-Leverage Plot

- This plot checks if there are any influential points, which could alter your analysis by excluding them

- The points that lie outside the dashed line, Cook's distance, are considered as influential points

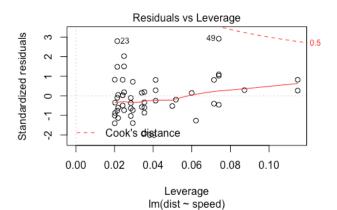
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Model Building Process

